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# Mendelian Randomization Analysis Reveals a Causal Effect of Urinary Sodium/Urinary Creatinine Ratio on Kidney Function in Europeans

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Salt restriction was recommended in clinical practice guideline for chronic kidney disease (CKD) treatment, but its effect on kidney outcomes remains conflicting. We aimed to test the causal effect of salt intake, using estimated 24-h sodium excretion from spot urinary sodium/urinary creatinine (UNa/UCr) ratio as a surrogate, on renal function using two-sample Mendelian randomization (MR). Genetic instruments for UNa/UCr were derived from a recent genome-wide association study of 218,450 European-descent individuals in the UK Biobank. Kidney outcomes were creatinine-based estimated glomerular filtration rate (eGFR<sub>crea</sub>) ( $N = 567,460$ ) and CKD (eGFR<sub>crea</sub> < 60 ml/min/1.73 m<sup>2</sup>,  $N$  cases = 41,395,  $N$  controls = 439,303) from the CKDGen consortium. Cystatin C-based eGFR (eGFR<sub>cys</sub>) and eGFR<sub>crea</sub> single-nucleotide polymorphisms associated with blood urea nitrogen (BUN) were used for sensitivity analyses. MR revealed a causal effect of UNa/UCr on higher eGFR<sub>crea</sub> [ $\beta = 0.14$ , unit change in log ml/min/1.73 m<sup>2</sup> per UNa/UCr ratio; 95% confidence interval (CI) = 0.07 – 0.20,  $P = 2.15 \times 10^{-5}$ ] and a protective effect against CKD risk (odds ratio = 0.24, 95% CI = 0.14 to 0.41,  $P = 1.20 \times 10^{-7}$ ). The MR findings were confirmed by MR-Egger regression, weighted median MR, and mode estimate MR, with less evidence of existence of horizontal pleiotropy. Consistent positive causal effect of UNa/UCr on eGFR<sub>cys</sub> was also detected. On the other hand, bidirectional MR suggested inconclusive results of CKD, eGFR<sub>crea</sub>, eGFR<sub>crea</sub> (BUN associated), and eGFR<sub>cys</sub> on UNa/UCr. The average 24-h sodium excretion was estimated to be approximately 2.6 g per day for women and 3.7 g per day for men. This study provides evidence that sodium excretion, well above the recommendation of <2 g per day of sodium intake, might not have a harmful effect on kidney function. Clinical trials are warranted to evaluate the sodium restriction target on kidney function.

**Keywords:** salt restriction, urinary sodium/urinary creatinine ratio, kidney function, Mendelian randomization, causal effect

## INTRODUCTION

Chronic kidney disease (CKD) accounts for 10% to 15% population worldwide, which is associated with high annual rates of mortality and cardiovascular complications (Xie et al., 2018). Identifying potentially modifiable risk factors for CKD is critical in order to devise effective, population-based preventive strategies.

A salt restriction <90 mmol per day (equivalent to 5 g per day of salt or 2 g per day of sodium) is recommended in the 2012 Kidney Disease: Improving Global Outcomes (KDIGO) clinical practice guideline (Kidney Disease: Improving Global Outcomes [KDIGO] and CKD Work Group, 2012). The Working Group graded this recommendation as “1C,” indicating that controlling salt intake can be evaluated as a candidate for developing a policy or a performance measure (Grade level 1), but from low quality of evidence (Grade C). This recommendation was mainly based on expert opinion and extrapolation from previous studies between sodium intake and blood pressure. Observational epidemiology studies have indicated that sodium intake was positively associated with blood pressure and proteinuria, which are strong predictors of CKD progression and cardiovascular diseases (CVDs) (He et al., 2013). In addition, existing clinical trials in sodium restriction showed significant reductions in blood pressures and proteinuria (Vogt et al., 2008; Suckling et al., 2010; Slagman et al., 2011; Graudal et al., 2012; de Brito-Ashurst et al., 2013). However, prospective cohort studies have reported conflicting findings on the association between dietary sodium intake and CVD (Cook et al., 2007; O'Donnell et al., 2011, 2014). Moreover, emerging studies suggested that extreme strict sodium control may be harmful rather than beneficial as previously recommended in CKD patients (Fan et al., 2014). Debates about optimal sodium intake in the general population are still ongoing (Mente et al., 2016, 2018). During the 2017 KDIGO controversies conference, stakeholders questioned whether a level 1C recommendation of sodium restriction (<2 g of sodium per day) was too strong based on the current evidence (Cheung et al., 2019).

Given that few randomized controlled trials have been conducted to infer causality between salt intake and kidney outcomes, the genetic variants associated with potential risk factors offer the opportunity to test the causal effect of salt intake on CKD through Mendelian randomization (MR) (Smith and Ebrahim, 2003; Zheng et al., 2017). In essence, MR exploits the random allocation of genetic variants at conception and therefore is less susceptible to confounding compared to traditional observational studies. Furthermore, MR studies give the effects of lifetime exposures and so could evaluate the long-term effect. Recently, Zanetti et al. performed genome-wide association studies (GWASs) of urinary electrolyte excretion, including the urinary sodium/urinary creatinine (UNa/UCr) ratio, in up to 327,616 unrelated individuals of European ancestry from the UK Biobank (Zanetti et al., 2019). Since spot urine samples were commonly used to estimate 24-h sodium excretion as a surrogate for salt intake (Lee et al., 2013; Mente et al., 2016, 2018), we aimed to assess the causal effect of salt intake, using UNa/UCr as proxy, on kidney function in Europeans using MR (Hemani et al., 2018).

## MATERIALS AND METHODS

### Genetic Predictors of Urinary Sodium Excretion

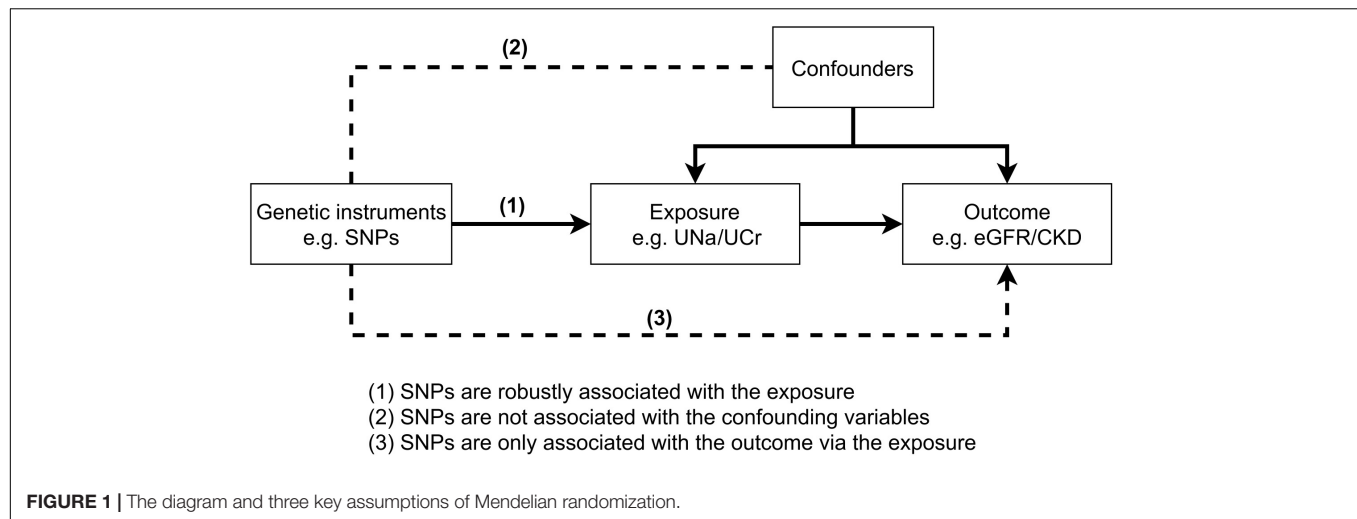
In a recent genetic study of urinary sodium excretion normalized for creatinine (calculated as UNa/UCr), 327,616 eligible individuals from the UK Biobank were included and two-thirds of the sample ( $n = 218,450$ ) were randomly selected for discovery GWAS (Zanetti et al., 2019). Association analysis was conducted using linear regression of UNa/UCr levels on imputed genotypes, assuming an additive model between phenotype and genotype dosages. The model was adjusted for age, sex, batch, and the first 10 genotype principal components. Six conditional distinguished genetic variants strongly associated with UNa/UCr ( $P < 5 \times 10^{-8}$ ) were selected as candidate instruments (**Supplementary Table S1**).

### Genetic Associations of Kidney Function Phenotypes

Creatinine-based estimated glomerular filtration rate (eGFR<sub>crea</sub>) and CKD (eGFR<sub>crea</sub> < 60 ml/min/1.73 m<sup>2</sup>, binary phenotype) were used as proxies for kidney function phenotypes. eGFR<sub>crea</sub> was estimated from serum creatinine using the Chronic Kidney Disease Epidemiology Collaboration (CKD-EPI) formula (Levey et al., 2009). It was estimated as  $141 \times \min(\text{Scr}/\kappa, 1)^\alpha \times \max(\text{Scr}/\kappa, 1)^{-1.209} \times 0.993^{\text{Age}} \times 1.018$  (if female)  $\times 1.159$  (if black), where Scr is the serum creatinine,  $\kappa$  is 0.7 for females and 0.9 for males,  $\alpha$  is -0.329 for females and -0.411 for males, min indicates the minimum of Scr/ $\kappa$  or 1, and max indicates the maximum of Scr/ $\kappa$  or 1. GWAS summary statistics of eGFR<sub>crea</sub> ( $N = 567,460$ ) and CKD ( $N$  cases = 41,395;  $N$  controls = 439,303;  $N$  total = 480,698) from the CKDGen consortium were utilized (Wuttke et al., 2019). The GWASs were imputed using the 1000 Genomes phase 1 v3 training set with 38 million variants. Genetic associations were adjusted for study specific covariates such as age and sex, and the first 10 principal components were also used as covariates in the association model of the GWAS to control for the population stratification.

### Mendelian Randomization

MR is an instrumental variable method that uses genetic variants as instruments to test the casual relationships between an exposure (e.g., UNa/UCr) and an outcome (e.g., eGFR<sub>crea</sub>/CKD). The causal association between the exposure and the outcome will be valid when the three core MR assumptions were satisfied (**Figure 1**). In our initial MR analysis, UNa/UCr was treated as exposure and eGFR<sub>crea</sub>/CKD as outcomes, using the UNa/UCr-associated single-nucleotide polymorphisms (SNPs) as instrument variables. We used the random-effect inverse variance weighted (IVW) method (Burgess et al., 2013) to estimate the MR effect, which pools the individual SNP Wald estimator (Lawlor et al., 2008) by meta-analysis. The Cochran's Q and Rucker's Q statistics were used to estimate the level of heterogeneity of the IVW and MR-Egger regression separately, in which the degree of freedom for Cochran's Q is



the number of SNPs minus one and that for Rucker's Q is the number of SNPs minus two.

We further tested the MR assumptions using the following sensitivity analyses. MR-Egger regression (Bowden et al., 2015) was used to test pleiotropic effects of the instruments on outcomes, independent of exposures. The MR-Egger regression intercept was recorded as an indicator of pleiotropy. We also performed MR analyses using the weighted median MR approach (Bowden et al., 2016) and the mode estimate MR approach (Hartwig et al., 2017), which provides consistent causal estimates of the exposure on the outcome even when up to 50% (or even 0%) of the information contributing to the analysis comes from genetic variants that exhibit pleiotropy. If all approaches (i.e., IVW MR, MR-Egger regression, weighted median MR, and mode estimate MR) provide consistent causal estimates of UNa/UCr on eGFRcrea/CKD, the MR findings will be more robust. To exclude the influence of pleiotropy instruments on the MR estimates, we applied MR-PRESSO (Verbanck et al., 2018) as another sensitivity analysis, which attempts to reduce heterogeneity in the estimate of the causal effect by removing SNPs that contribute to the heterogeneity disproportionately more than expected.

## Directionality Analysis

### Steiger Filtering

When applying MR, we assume that SNPs used as proxies of UNa/UCr exert their primary association on UNa/UCr and that any correlation with eGFRcrea or CKD is a consequence of a causal effect of UNa/UCr on eGFRcrea or CKD. We therefore performed Steiger filtering analyses (Hemani et al., 2018), implemented in the TwoSampleMR R package, to test the directionality of UNa/UCr instruments. The underline concept of Steiger filtering is a simple inequality. Given that phenotype A causes phenotype B, then we would expect that

$$\sum_{i=1}^M \text{cor}(g_i, A)^2 > \sum_{i=1}^M \text{cor}(g_i, B)^2$$

since  $\text{cor}(g_i, B)^2 = \text{cor}(A, B)^2 * \text{cor}(g_i, A)^2$ , where “cor” denotes correlation, and the vector  $g$  contains a set of SNPs that influence phenotype A. For any instrument that had a  $\text{cor}(g_i, A)^2 < \text{cor}(g_i, B)^2$ , they showed evidence of primarily affecting eGFR rather than UNa/UCr. We thus removed these UNa/UCr instruments and performed the MR analyses using the remaining instruments.

### Bidirectional MR

To investigate the potential reverse causality of eGFRcrea/CKD on UNa/UCr, we conducted a bidirectional MR. We first found 256 conditional independent variants associated with eGFRcrea and 23 independent variants associated with CKD (Supplementary Table S2) using the linkage disequilibrium clumping ( $r^2 < 0.001$ ) function implemented in the TwoSampleMR R package. We then used eGFRcrea or CKD as exposures and UNa/UCr as the outcome for the bidirectional MR. Considering that eGFRcrea has GFR and marker determinants, we further performed sensitivity analyses by using blood urea nitrogen (BUN) GWAS (Wuttke et al., 2019) to prioritize eGFRcrea-associated SNPs as exposures. More specifically, we required the eGFRcrea SNPs be associated with BUN in the opposite direction (since higher GFR would lead to lower BUN) and the eGFRcrea SNPs be associated with BUN with a Bonferroni-corrected significance. In addition, cystatin C, an alternative biomarker of kidney function, was measured in a subset of European ancestry participants from the CKDGen consortium ( $N = 32,861$ ) (Li et al., 2017). Cystatin C-based eGFR (eGFRcys) could be estimated as  $76.7 \times \text{serum cystatin C}$  (Levey et al., 2011). To further exclude the possible bias caused by creatinine, we extracted five SNPs that robustly associated with eGFRcys (Li et al., 2017) and used them as instruments for a sensitivity MR analysis against UNa/UCr.

### MR and Visualization Software

The MR results were presented as scatter plots and forest plots using a code derived from the ggplot2 package in R<sup>1</sup>. All MR and sensitivity analyses were conducted using the

<sup>1</sup><https://cran.r-project.org/web/packages/ggplot2/index.html>



TwoSampleMR R package<sup>2</sup> (Hemani et al., 2018). MR-PRESSO test was conducted by using the MR-PRESSO R package<sup>3</sup>. The number of distributions was set to 10,000, and the threshold was set to 0.05.

## 24-h Urinary Sodium Estimation

The exact intake of sodium cannot be measured precisely but can be estimated from 24-h sodium excretion (Kawano et al., 2007). Thus, to estimate the sodium intake, we used the INTERSALT (International Cooperative Study on Salt, Other Factors, and Blood Pressure) equation (Brown et al., 2013), developed in a sample of 3,093 European-descent individuals aged 20–59 years, to estimate the 24-h urine sodium excretion based on the reported spot urine measurements. The equation was described as follows.

For men:

24-h sodium excretion (mg) =  $23 \times \{25.46 + [0.46 \times \text{spot urine sodium (mmol/L)}] - [2.75 \times \text{spot urine creatinine (mmol/L)}] - [0.13 \times \text{spot urine potassium (mmol/L)}] + [4.10 \times \text{body mass index (kg/m}^2\text{)}] + [0.26 \times \text{age (years)}]\}$ .

For women:

24-h sodium excretion (mg) =  $23 \times \{5.07 + [0.34 \times \text{spot urine sodium (mmol/L)}] - [2.16 \times \text{spot urine creatinine (mmol/L)}] - [0.09 \times \text{spot urine potassium (mmol/L)}] + [2.39 \times \text{body mass index (kg/m}^2\text{)}] + [2.35 \times \text{age (years)}] - [0.03 \times \text{age}^2 \text{ (years)}]\}$ .

## RESULTS

### Effect of UNa/UCr on eGFRcrea and CKD

Figure 2 showed the MR results of UNa/UCr on eGFRcrea and CKD. For eGFRcrea, the IVW MR suggested that UNa/UCr was positively associated with eGFRcrea [ $\beta = 0.14$ , unit change in log ml/min/1.73 m<sup>2</sup> per UNa/UCr ratio; 95% confidence interval (CI) = 0.07 – 0.20,  $P = 2.15 \times 10^{-5}$ , Figure 2A]. Similarly, the IVW MR suggested that increased UNa/UCr was associated with reduced CKD risk [odds ratio (OR) = 0.24, 95% CI = 0.14 to 0.41,  $P = 1.20 \times 10^{-7}$ , Figure 2B]. The SNP MR results presented by the forest plots suggested that the causal effects of UNa/UCr on eGFRcrea/CKD were similar across instruments, and the observed MR findings were not massively influenced by any single instrument (Figures 2A,B). Figures 2C,D showed scatter plots of the SNP-outcome associations against the SNP-UNa/UCr associations, allowing visualization of the causal effect for individual SNP on eGFRcrea/CKD.

Cochrane's Q test ( $P = 7.03 \times 10^{-7}$  and  $5.83 \times 10^{-3}$  for eGFRcrea and CKD as outcomes, respectively) and Rocker's Q-test ( $P = 6.72 \times 10^{-7}$  and  $2.29 \times 10^{-3}$  for eGFRcrea and CKD as outcomes, respectively) suggested the presence of considerable heterogeneity in the analysis. However, the MR-Egger intercept suggested little evidence of horizontal pleiotropy ( $P = 0.65$  and 0.98 for eGFRcrea and CKD as outcome, respectively) (Supplementary Table S3).

<sup>2</sup><http://www.mrbase.org>

<sup>3</sup><https://github.com/rondolab/MR-PRESSO>

## Sensitivity Analyses

Table 1 showed the MR estimates of UNa/UCr on eGFRcrea and CKD using other sensitivity MR approaches, which are more robust to pleiotropy. For eGFRcrea, the IVW MR result was consistent with the casual estimates using MR-Egger regression ( $\beta = 0.32$ , 95% CI = 0.06 – 0.58,  $P = 1.66 \times 10^{-2}$ ), weighted median MR ( $\beta = 0.07$ , 95% CI = 0.04 – 0.11,  $P = 8.68 \times 10^{-6}$ ), and mode-based estimate MR approaches ( $\beta = 0.05$ , 95% CI = 0.02 – 0.09,  $P = 1.82 \times 10^{-2}$ ), which increased the robustness of the MR results. For CKD, the IVW MR estimates were concordant in direction with the MR estimates using MR-Egger (OR = 0.32, 95% CI = 0.02 – 4.97,  $P = 0.46$ ), weighted median MR (OR = 0.23, 95% CI = 0.14 – 0.40,  $P = 1.40 \times 10^{-7}$ ), and mode-based MR approaches (OR = 0.20, 95% CI = 0.10 – 0.41,  $P = 7.55 \times 10^{-3}$ ). In consistent with these results, we also observed positive causal effect of UNa/UCr on eGFRcys (Supplementary Table S4). Given that some of the instruments might be influenced by horizontal pleiotropy, we applied the MR-PRESSO outlier test and identified three UNa/UCr instruments as potential outliers. The MR estimate after removing the outliers identified by MR-PRESSO suggested a consistent positive effect of UNa/UCr on eGFRcrea (MR-PRESSO estimate:  $\beta = 0.129$ , 95% CI = 0.076 – 0.182,  $P = 4.14 \times 10^{-2}$ ; Supplementary Table S5).

## Directionality Tests

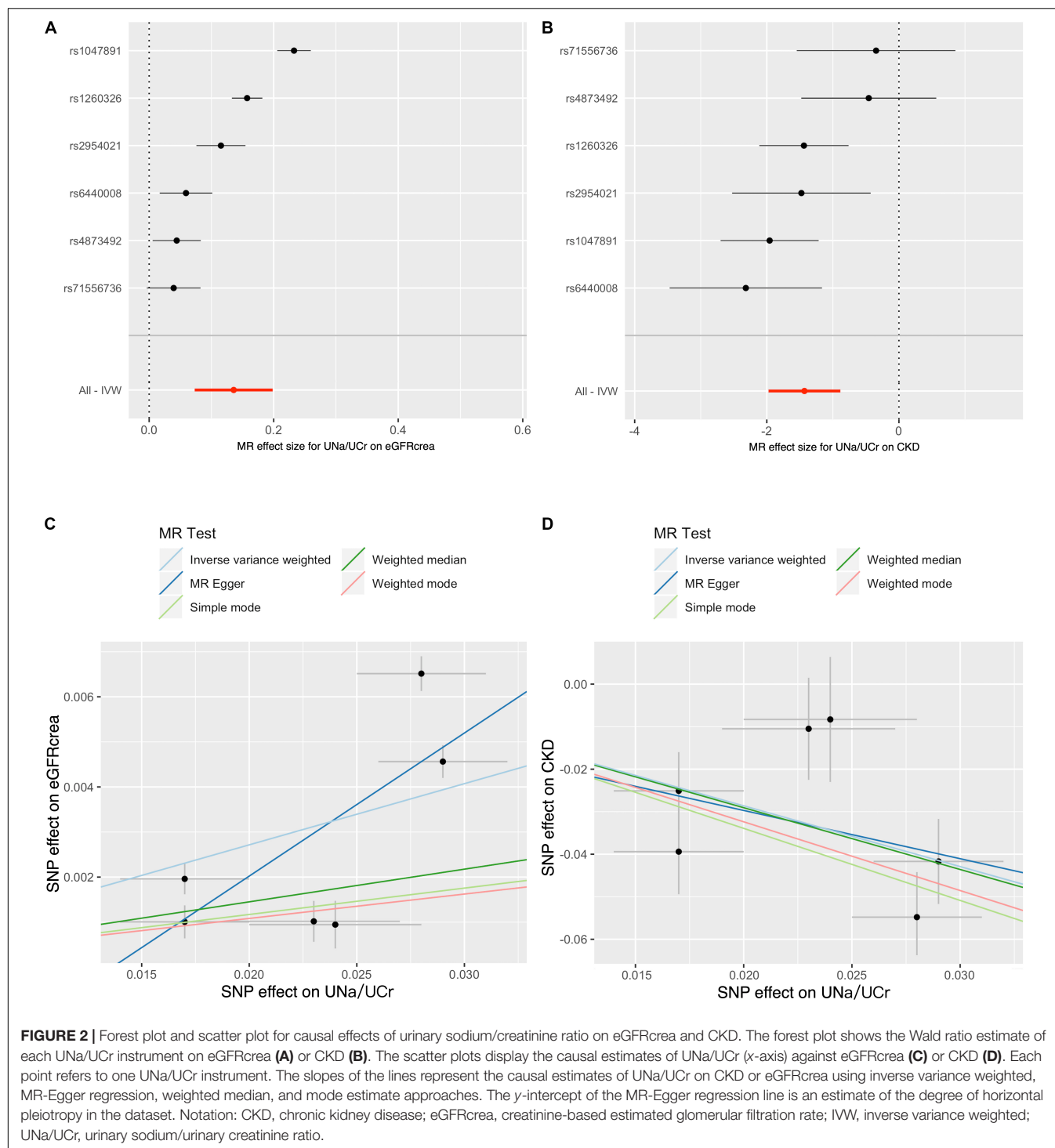
The Steiger filtering results suggested that all instruments were primarily influencing UNa/UCr than affecting eGFRcrea as a consequence (Supplementary Table S6), which further confirmed the robustness of the instruments. As shown in Figure 3, bidirectional MR using CKD, eGFRcrea, eGFRcrea (BUN associated), and eGFRcys as exposures and UNa/UCr as an outcome suggested inconsistent evidence of causal effect of kidney function on UNa/UCr (Supplementary Table S7).

## Sodium Intake Estimation

The mean age, body mass index, UNa/UCr, and urinary potassium/UCr were 56.88, 27.39, 10.60, and 8.48, respectively, in the discovery population and were 56.87, 27.42, 10.62, and 8.44, respectively, in the replication population (Zanetti et al., 2019). Using these baseline characteristics, the average 24-h sodium excretion was estimated to be approximately 2.6 g per day for women and 3.7 g per day for men, well above the recommended restriction of <2 g per day of sodium intake (Supplementary Table S8).

## DISCUSSION

Sodium intake was reported to be one of the most important modifiable risk factors influencing blood pressure and proteinuria, yet its impact on renal outcomes is uncertain. Using a two-sample MR approach, we tested the causal effect of sodium intake, using estimated 24-h sodium excretion from a spot UNa/UCr ratio as a surrogate, on kidney function. We found that each unit increase in the UNa/UCr ratio was associated with a 0.14 unit increase in log(eGFRcrea) and with reduced CKD risk (OR = 0.24). These MR findings were confirmed by



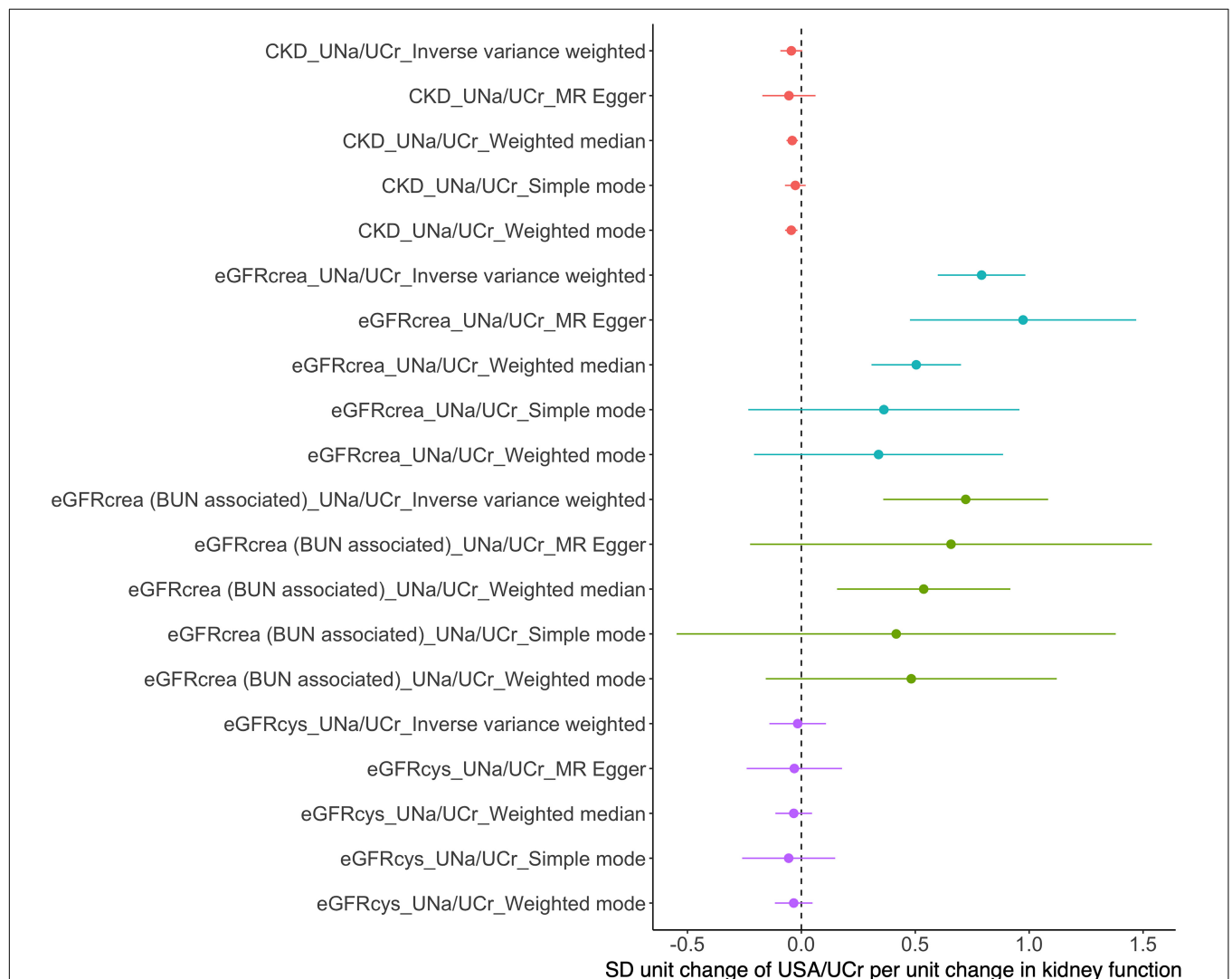
MR-Egger regression, weighted median MR, and mode estimate MR, which suggested less evidence of existence of horizontal pleiotropy. Steiger filtering suggested less evidence of reverse causality of CKD on UNa/UCr. Cochran's *Q* and Rucker's *Q* statistics suggested the presence of considerable heterogeneity in the analysis. However, the MR-Egger intercept suggested little evidence of horizontal pleiotropy. Consistent positive causal

effect of UNa/UCr on eGFRcrea was also detected. Bidirectional MR using CKD, eGFRcrea, eGFRcrea (BUN associated), and eGFRcrea as exposures showed weak evidence supporting a causal relationship for kidney function on UNa/UCr. The average 24-h sodium excretion was estimated to be approximately 2.6 g per day for women and 3.7 g per day for men, well above the recommendation of <2 g per day of sodium intake

**TABLE 1** | Mendelian randomization analyses of causal effects of urinary sodium/urinary creatinine ratio on eGFR<sub>crea</sub> and CKD.

Trait	Number of SNPs	Methods	Estimator (95% CI)	P
eGFR <sub>crea</sub>	6	MR IVW	$\beta = 0.14$ (0.07 – 0.20)	$2.15 \times 10^{-5}$
	6	MR-egger	$\beta = 0.32$ (0.06 – 0.58)	$1.66 \times 10^{-2}$
	6	MR-weighted median	$\beta = 0.07$ (0.04 – 0.11)	$8.68 \times 10^{-6}$
	6	MR-weighted mode	$\beta = 0.05$ (0.02 – 0.09)	$1.82 \times 10^{-2}$
CKD	6	MR IVW	OR = 0.24 (0.14 – 0.41)	$1.20 \times 10^{-7}$
	6	MR-egger	OR = 0.32 (0.02 – 4.97)	0.46
	6	MR-weighted median	OR = 0.23 (0.14 – 0.40)	$1.40 \times 10^{-7}$
	6	MR-weighted mode	OR = 0.20 (0.10 – 0.41)	$7.55 \times 10^{-3}$

The confidence interval of MR-Egger estimates was wider compared to that of other methods because the statistical power of the MR-Egger was, in general, lower than that of MR IVW and weighted median. CKD, chronic kidney disease; eGFR<sub>crea</sub>, creatinine-based estimated glomerular filtration rate; MR, Mendelian randomization; IVW, inverse variance weighted, OR, odds ratio; SNP, single-nucleotide polymorphism.



**FIGURE 3** | Mendelian randomization analyses of causal effects of eGFR/CKD on urinary sodium/creatinine ratio. BUN, blood urea nitrogen; eGFR<sub>crea</sub>, creatinine-based estimated glomerular filtration rate; eGFR<sub>crea</sub> (BUN associated), the eGFR<sub>crea</sub>-associated SNPs were required to be associated with BUN in opposite direction (since higher GFR would lead to lower BUN) and to be associated with BUN with a Bonferroni-corrected significance; eGFR<sub>cys</sub>, cystatin C-based estimated glomerular filtration rate.

in the KDIGO guideline. Our results support that reducing dietary sodium intake under 2 g per day in CKD might be overcontrolled and may not have a beneficial effect on reducing CKD risk as expected.

Sodium intake was regarded as one of the important modifiable risk factors for CKD progression. The exact intake of sodium cannot be measured precisely but can be estimated from dietary questionnaires or 24-h sodium excretion (Kawano et al., 2007). However, questionnaires are usually self-reported and require calculations by a nutritionist and thus may be imperfect measures of sodium intake (Leiba et al., 2005), while the 24-h urine sample collection is limited to poor compliance and cost. Thus, studies usually used spot urine samples to estimate 24-h sodium excretion as a surrogate for sodium intake (Mente et al., 2016, 2018). To compensate for variation in urine concentration, urinary sodium corrected by urinary creatinine could be used for assessing sodium excretion in urine over 24-h (Lee et al., 2013). Using these methods, previous observational studies have provided inconsistent results on the effects of dietary sodium intake on the risk of CKD. Some studies suggested that urinary sodium excretion was associated with increased risk of CKD progression and incidence of end-stage renal disease (ESRD) (McCausland et al., 2012; Vegter et al., 2012; Ohta et al., 2013; McQuarrie et al., 2014), while others indicated that urinary sodium excretion was not associated with the risk of CKD or renal failure (Dunkler et al., 2013; Fan et al., 2014; Smyth et al., 2014), even protected against cumulative incidence of ESRD (Ekinci et al., 2011; Thomas et al., 2011). The inconsistent results may be due to the fact that these studies were observational studies that cannot make causal inference.

Until now, previous trials mainly focused on the blood pressure lowering effect of sodium control, but there are limited evidence detailing the effects of sodium restriction in CKD patients. For example, the Holland Nephrology Study (HONEST) suggested that the addition of a low-sodium diet further reduced 11 mmHg systolic blood pressure in patients with treatment with an angiotensin converting enzyme inhibitor (Slagman et al., 2011). Vogt found reductions of 6/3 mmHg systolic blood pressure/diastolic blood pressure in proteinuric patients without diabetes who received a low-sodium diet (Vogt et al., 2008). Two more recent studies provided new clinical trial data from randomized, double-blind, crossover trials in patients with stages 3 and 4 CKD (McMahon et al., 2013; Saran et al., 2017). They found that salt restriction resulted in statistically significant reductions of blood pressure and albuminuria. However, although salt intake has been extensively suggested to be associated with blood pressure, high sodium intake-related increase in blood pressure might not entirely explain the elevated risk of CKD progression and renal failure in patients with CKD. What's more, it must be noted that it is unable to compare the clinical outcomes among patients with urinary sodium excretion <2 g per day recommended in clinical practice guideline, which appears difficult to achieve. Results from most trials showed that only high urinary sodium excretion (approximately 150–200 mmol or 3.5–4.5 g sodium per day), well above the recommended limit (2 g sodium per day), was associated with increased risk of CKD progression

or all-cause mortality (Norris et al., 2006; Lambers et al., 2012; McMahon et al., 2013; He et al., 2016). Thus, using these data is hard to assess whether sodium restriction <2 g per day is suggestively hazardous or intuitively advantageous (Anderson and Ix, 2013).

In the current study, we used the two-sample MR approach to test the causal effect of UNa/UCr on kidney outcomes. We further estimated the 24-h urine sodium excretion to estimate sodium intake using the previously published sodium excretion estimating equation based on spot urine measurements, which was developed and validated in European populations (Brown et al., 2013). We found that UNa/UCr showed causal effect on higher eGFR<sub>crea</sub> and lower CKD risk, while weak evidence of causal association between kidney function and UNa/UCr was observed. Using baseline characteristics of the UK Biobank participants, the average 24-h sodium excretion was estimated to be approximately 2.6 g per day for women and 3.7 g per day for men. Previous studies suggested that this estimation equation might underestimate the urinary sodium excretion (Allen et al., 2017). Although potential bias exists in estimating 24-h sodium excretion using casual urinary test (Dougher et al., 2016), our results support that reducing dietary sodium intake under 2 g per day in CKD was too restrictive and might not be beneficial to reduce the risk of CKD development or progression as expected.

Several limitations should be mentioned in this study. First, 24-h urine collection was estimated rather than performed in the study population; target levels of UNa/UCr for protecting effects were not evaluated. In the future, additional large randomized controlled trials are required to determine the optimal target sodium intake. Second, we currently have limited access to the individual level genotype and phenotype data to perform further stratification of groups by instruments or outcomes and distributions. This type of analysis is needed to determine the optimal target of sodium intake in CKD in the future. Third, this study is limited to the only six instruments of UNa/UCr, even though they were derived from the well-powered and most recent UNa/UCr GWAS. Most of the current MR sensitivity analysis methods such as MR-Egger regression and weighted median need a large number of independent instrumental SNPs in order to test for pleiotropy. Thus, the MR results presented in this paper should be taken with caution. Fourth, the MR Steiger filtering approach has limitations, i.e., some levels of horizontal pleiotropy, where the SNP influences the outcome through some pathways other than the exposure, could induce problems because this is a means by which the instrument is invalid. Besides, some levels of unmeasured confounding between the exposure and the outcome could lead to inference of the wrong causal direction. We therefore applied a well-validated method MR-PRESSO. After controlling for pleiotropy instruments, our finding was still consistent. Finally, as in most epidemiological studies, MR assumes a linear relation between UNa/UCr and eGFR/CKD, which might not invariably be the case.

In conclusion, we found that a higher UNa/UCr ratio, which reflects proportionally higher salt intake, showed causal effects on higher eGFR and lower CKD risk. Further long-term studies to determine optimal target of sodium intake in CKD are required.



## DATA AVAILABILITY STATEMENT

The kidney function datasets can be found from <http://ckdgen.imbi.uni-freiburg.de/>. The summary statistics of UNa/UCr was available at GRASP resource (<https://grasp.nhlbi.nih.gov/FullResults.aspx>, PMID: 30910378).

## AUTHOR CONTRIBUTIONS

YZ, JZ, and HZ contributed to the conception and design of the study. YZ prepared the data and wrote the draft of the manuscript. JZ performed most of the statistical analysis. All authors contributed to manuscript revision and read and approved the submitted version.

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## SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: <https://www.frontiersin.org/articles/10.3389/fbioe.2020.00662/full#supplementary-material>

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**Conflict of Interest:** The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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